

A Strong Limit Theorem for Sequences of Blockwise and Pairwise Negative Quadrant M -Dependent Random Variables

VU THI NGOC ANH

Department of Mathematics, Hoa Lu University, Ninh Binh, Vietnam
anhyk86@gmail.com

Abstract. In this paper, we establish a Marcinkiewicz-Zygmund type strong law for sequences of blockwise and pairwise negative quadrant m -dependent random variables. The sharpness of the results is illustrated by an example.

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1. Introduction

The concept of negative quadrant dependence was introduced by Lehmann [1]. The concept of blockwise m -dependence and blockwise quasiorthogonality for a sequence of random variables was introduced by Móricz [3]. The strong laws for blockwise independence case or blockwise orthogonal case then was studied by some authors. We refer to Rosalsky and Thanh [7], Quang and Thanh [5] for Banach spaces valued case and Quang and Thanh [6], Thanh [10] for multi-dimension case. Thanh and Anh [11] established a strong law of large numbers for blockwise and pairwise m -dependent random variables which extends the result of Thanh [8] to the arbitrary blocks case and also provided an example to illustrate the main result. In Thanh and Anh [11], authors considered a sequence of random variables which is blockwise and pairwise m -dependent with respect to the arbitrary blocks.

In this note, we consider a sequence of blockwise and pairwise negative quadrant m -dependent random variables $\{X_n, n \geq 1\}$ which is stochastically dominated by a random variable X . We establish a Marcinkiewicz-Zygmund type strong law of large numbers which extends the result of Thanh and Anh [11] to the blockwise and pairwise negative quadrant m -dependent case. We also provide an example to illustrate the main result.

Let X and Y be random variables. We say that X and Y are *negative quadrant dependent* if

$$P(X \leq x, Y \leq y) \leq P(X \leq x)P(Y \leq y), \quad \forall x, y \in \mathbb{R}.$$

A sequence of random variables $\{X_n, n \geq 1\}$ is said to be *pairwise negative quadrant dependent* if for all $i \neq j$, X_i and X_j are negative quadrant dependent.

Let m be a fixed nonnegative integer. We say that a collection $\{X_j, 1 \leq j \leq n\}$ of n random variables is *pairwise negative quadrant m -dependent* if either $n \leq m + 1$ or $n > m + 1$ and X_i and X_j are negative quadrant dependent whenever $j - i > m$.

Let $\{\beta_k, k \geq 1\}$ be a strictly increasing sequence of positive integers with $\beta_1 = 1$ and set $B_k = [\beta_k, \beta_{k+1})$.

A sequence of random variables $\{X_n, n \geq 1\}$ is said to be *blockwise and pairwise negative quadrant m -dependent* with respect to the blocks $\{B_k, k \geq 1\}$ if for each $k \geq 1$, the random variables $\{X_i, i \in B_k\}$ are pairwise negative quadrant m -dependent.

For $\{\beta_k, k \geq 1\}$ and $\{B_k, k \geq 1\}$ as above, we introduce the following notation:

$$B^{(l)} = \{k : 2^l \leq k < 2^{l+1}\}, l \geq 0,$$

$$B_k^{(l)} = B_k \cap B^{(l)}, k \geq 1, l \geq 0,$$

$$I_l = \{k \geq 1 : B_k^{(l)} \neq \emptyset\}, l \geq 0,$$

$$r_k^{(l)} = \min\{r : r \in B_k^{(l)}\}, k \in I_l, l \geq 0,$$

$$c_l = \text{card}I_l, l \geq 0,$$

$$\varphi(n) = \sum_{l=0}^{\infty} c_l I_{B^{(l)}}(n), n \geq 1,$$

$$\psi(n) = \max_{k \leq n} \varphi(k), n \geq 1$$

where $I_{B^{(l)}}$ denotes the indicator function of the set $B^{(l)}$, $l \geq 0$.

Random variables $\{X_n, n \geq 1\}$ are said to be a *stochastically dominated* by random variable X if for some constant $C < \infty$

$$P(|X_n| > t) \leq CP(|X| > t), \forall t \geq 0, \forall n \geq 1.$$

2. Main result

Throughout this section, the logarithms are to the base 2, the symbol C denotes a generic constant ($0 < C < \infty$) which is not necessarily the same one in each appearance.

Before establishing main result, we state two lemmas. The first lemma can be obtained by using a method similar to that used in the proof the Rademacher-Menshov inequality and Lemma 2.2 of Li, Rosalsky and Volodin [2].

Lemma 2.1. *If $\{X_n, n \geq 1\}$ is a sequence of pairwise negative quadrant dependent mean 0 random variables, then*

$$E \left(\max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| \right)^2 \leq C(\log 4n)^2 \sum_{j=1}^n EX_j^2.$$

The second lemma can be obtained by using a method similar to that used in the proof the Lemma 3 of Thanh [8] and Lemma 2.2. It extends the Lemma 3 of Thanh [8] to the blockwise and pairwise negative quadrant m -dependent case.

Lemma 2.2. *If $\{X_j, 1 \leq j \leq n\}$ is a collection of pairwise negative quadrant m -dependent mean 0 random variables, then*

$$E \left(\max_{1 \leq k \leq n} \left| \sum_{j=1}^k X_j \right| \right)^2 \leq C(m+1)(\log 4n)^2 \sum_{j=1}^n EX_j^2.$$

With the preliminaries accounted for, the main result may be established.

Theorem 2.1. *Let $1 \leq r < 2$ and $\{X_n, n \geq 1\}$ be a sequence of random variables which is blockwise and pairwise negative quadrant m -dependent with respect to the blocks $\{B_k, k \geq 1\}$. Suppose that $\{X_n, n \geq 1\}$ is stochastically dominated by a random variable X . If*

$$(2.1) \quad E(|X|^r (\log^+ |X|)^2) < \infty,$$

then

$$(2.2) \quad \lim_{n \rightarrow \infty} \frac{1}{n^{1/r} \psi^{1/2}(n)} \sum_{j=1}^n (X_j - EX_j) = 0 \text{ a.s.}$$

Proof. Set

$$\begin{aligned} Y_n &= X_n I(|X_n| \leq n^{1/r}) + n^{1/r} I(X_n > n^{1/r}) - n^{1/r} I(X_n < -n^{1/r}), \\ Y_n^{(+)} &= X_n^+ I(X_n \leq n^{1/r}) + n^{1/r} I(X_n > n^{1/r}), \\ Y_n^{(-)} &= X_n^- I(X_n \geq -n^{1/r}) + n^{1/r} I(X_n < -n^{1/r}), \quad n \geq 1 \end{aligned}$$

and

$$\begin{aligned} T_{k(l)}^{(+)} &= \max_{j \in B_k^{(l)}} \left| \sum_{i=r_k^{(l)}}^j (Y_i^{(+)} - EY_i^{(+)}) \right|, \quad k \in I_l, \quad l \geq 0, \\ \tau_l^{(+)} &= \frac{1}{\left(2^{\frac{l+1}{r}} - 2^{\frac{l}{r}}\right) \psi^{\frac{1}{2}}(2^l)} \sum_{k \in I_l} T_{k(l)}^{(+)}, \quad l \geq 0. \end{aligned}$$

It follows from Lemma 2.1 of Li, Rosalsky and Volodin [2] that $\{Y_n^{(+)}, n \geq 1\}$ and $\{Y_n^{(-)}, n \geq 1\}$ are sequences of random variables which are blockwise and pairwise negative quadrant m -dependent with respect to the blocks $\{B_k, k \geq 1\}$.

Note at the outset that,

$$\begin{aligned} E(Y_n^{(+)})^2 &\leq 2 \int_0^{n^{\frac{1}{r}}} x P(|X_n| > x) dx, \\ E|X_n - Y_n| &\leq C \left(n^{\frac{1}{r}} P(|X_n| > n^{\frac{1}{r}}) + \int_{n^{\frac{1}{r}}}^{\infty} P(|X_n| > x) dx \right), \quad n \geq 1 \end{aligned}$$

and by using a method similar to that used in the proof of Theorem 1 of Thanh [8], we obtain

$$(2.3) \quad \sum_{n=1}^{\infty} \frac{\log^2 n}{n^{2/r}} E(Y_n^{(+)})^2 \leq CE(|X|^r (\log^+ |X|)^2) < \infty$$

and

$$(2.4) \quad \sum_{n=1}^{\infty} \frac{1}{n^{1/r}} E|X_n - Y_n| \leq CE(|X|^r \log^+ |X|) < \infty.$$

Note that for $l \geq 0$,

$$\begin{aligned} E(\tau_l^{(+)})^2 &\leq C \frac{1}{2^{\frac{2(l+1)}{r}} \psi(2^l)} c_l \sum_{k \in I_l} E(T_{k(l)}^{(+)})^2 \\ &\leq C \frac{1}{2^{\frac{2(l+1)}{r}}} \sum_{k \in I_l} (\log(4 \text{card} B_k^{(l)}))^2 \sum_{i \in B_k^{(l)}} E|Y_i^{(+)} - EY_i^{(+)}|^2 \\ &\quad \text{(by Lemma 2.3)} \\ &\leq C \frac{1}{2^{\frac{2(l+1)}{r}}} (\log 2^{l+2})^2 \sum_{i=2^l}^{2^{l+1}-1} E|Y_i^{(+)} - EY_i^{(+)}|^2 \\ &\leq C \sum_{i=2^l}^{2^{l+1}-1} \frac{(\log 4i)^2}{i^{\frac{2}{r}}} E(Y_i^{(+)})^2. \end{aligned}$$

It follows from (2.3) that $\sum_{l=0}^{\infty} E(\tau_l^{(+)})^2 < \infty$ and so by the Markov inequality and the Borel-Cantelli lemma ensures that

$$(2.5) \quad \lim_{l \rightarrow \infty} \tau_l^{(+)} = 0 \text{ a.s.}$$

Note that for $n \geq 1$, letting $M \geq 0$ be such that $2^M \leq n < 2^{M+1}$,

$$\begin{aligned} \frac{|\sum_{i=1}^n (Y_i^{(+)} - EY_i^{(+)})|}{n^{\frac{1}{r}} \psi^{\frac{1}{2}}(n)} &\leq \frac{\sum_{l=0}^M \sum_{k \in I_l} T_{k(l)}^{(+)}}{2^{\frac{M}{r}} \psi^{\frac{1}{2}}(2^M)} \\ &\leq \sum_{l=0}^M \frac{2^{\frac{l+1}{r}} - 2^{\frac{l}{r}}}{2^{\frac{M}{r}}} \tau_l^{(+)} \end{aligned}$$

and so (2.5) and Toeplitz lemma ensures that

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{r}} \psi^{\frac{1}{2}}(n)} \sum_{i=1}^n (Y_i^{(+)} - EY_i^{(+)}) = 0 \text{ a.s.}$$

Similarly,

$$\lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{r}} \psi^{\frac{1}{2}}(n)} \sum_{i=1}^n (Y_i^{(-)} - EY_i^{(-)}) = 0 \text{ a.s.}$$

and so $Y_n = Y_n^{(+)} - Y_n^{(-)}$, $n \geq 1$, we get

$$(2.6) \quad \lim_{n \rightarrow \infty} \frac{1}{n^{\frac{1}{r}} \psi^{\frac{1}{2}}(n)} \sum_{i=1}^n (Y_i - EY_i) = 0 \text{ a.s.}$$

By (2.4), (2.6) and by using a method similar to that used in the proof of Theorem 2.1 of Thanh and Anh [11], we obtain (2.2). ■

Note that if $\{X_n, n \geq 1\}$ is blockwise and pairwise m -dependent with respect to the blocks $\{B_k, k \geq 1\}$, then $\{X_n, n \geq 1\}$ is blockwise and pairwise negative quadrant m -dependent with respect to the blocks $\{B_k, k \geq 1\}$. So we get the following corollary which is the main result of Thanh and Anh [11].

Corollary 2.1. *Let $1 \leq r < 2$ and $\{X_n, n \geq 1\}$ be a sequence of random variables which is blockwise and pairwise m -dependent with respect to the blocks $\{B_k, k \geq 1\}$ and if (2.1) is satisfied, then (2.2) holds.*

Note that if $\beta_k = [q^{k-1}]$ for all large k and $q > 1$, then $c_l = \mathcal{O}(1)$, $\psi(n) = \mathcal{O}(1)$. So we get the following corollary.

Corollary 2.2. *Let $\{X_n, n \geq 1\}$ be a sequence of blockwise and pairwise negative quadrant m -dependent random variables with respect to the blocks $\{[2^{k-1}, 2^k], k \geq 1\}$ (or, more generally, with respect to the blocks $\{[\beta_k, \beta_{k+1}], k \geq 1\}$ where $\beta_k = [q^{k-1}]$ for all large k and $q > 1$) and if (2.1) is satisfied, then*

$$(2.7) \quad \lim_{n \rightarrow \infty} \frac{1}{n^{1/r}} \sum_{j=1}^n (X_j - EX_j) = 0 \text{ a.s.}$$

The following example is a modify of Example 2.6 in Thanh and Anh [11]. However, we try to construct with large blocks.

Example 2.1. Let $\{Y_n, n \geq 1\}$ be a sequence of 0-dependent identically distributed of $N(0, 1)$ random variables and let $3/2 \leq r < 2$. Let

$$X_n = Y_{n-k^3+1}, k^3 \leq n < (k+1)^3, k \geq 1.$$

Then $\{X_n, n \geq 1\}$ is blockwise and pairwise negative quadrant 0-dependent with respect to the blocks $\{[k^3, (k+1)^3], k \geq 1\}$ and (2.1) is satisfied, but $\{X_n, n \geq 1\}$ is not blockwise and pairwise negative quadrant m -dependent with respect to the blocks $\{[2^k, 2^{(k+1)}], k \geq 0\}$ for any non-negative integer m . Now, by noting that in $\beta_k = k^3, k \geq 1$ case $\psi(n) = \mathcal{O}(n^{1/3})$, so that by Corollary 2.4 we have

$$\lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n X_i}{n^{1/r+1/6}} = 0 \text{ a.s.}$$

Now, for $n = (M+1)^3 - 1$, we have

$$\begin{aligned} & \frac{\sum_{i=1}^n X_i}{n^{1/r}} \\ &= \frac{MY_1 + M(Y_2 + \cdots + Y_7) + \cdots + (Y_{M^3 - (M-1)^3 + 1} + \cdots + Y_{(M+1)^3 - M^3})}{((M+1)^3 - 1)^{1/r}} \\ &= S_{(M+1)^3 - 1}, \end{aligned}$$

where

$$S_{(M+1)^3 - 1} \sim N\left(0, \frac{M^4 + 4M^3 + 7M^2 + 2M}{2((M+1)^3 - 1)^{2/r}}\right),$$

so that (2.7) fails since $r \geq 3/2$.

Remark 2.1. Sequence $\{X_n, n \geq 1\}$ of random variables in Example 2.6 of Thanh and Anh [11] also is not blockwise and pairwise negative quadrant m -dependent with respect to the

blocks $\{[2^{k-1}, 2^k), k \geq 1\}$ for any non-negative integer m and (2.1) is satisfied but (2.7) fails. So it also shows that Theorem 2.3 is sharp. More precisely, it shows that for all $\varepsilon > 0$,

$$\limsup_{n \rightarrow \infty} \frac{|\sum_{i=1}^n X_i|}{n^{1/r-\varepsilon} \psi^{1/2}(n)} = \infty \text{ a.s.}$$

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